

## SPATIAL VARIABILITY OF SOIL NUTRIENTS USING GEOSPATIAL TECHNIQUES: A CASE STUDY IN SOILS OF SANWER TEHSIL OF INDORE DISTRICT OF MADHYA PRADESH

G.S. Tagore<sup>a</sup>, G.D. Bairagi<sup>\*b</sup>, R. Sharma<sup>b</sup> and P. K. Verma<sup>b</sup>

(\*bairagigd@gmail.com)

<sup>a</sup>Department of Soil Science and Agril. Chemistry, JNKVV, Jabalpur (M.P.)

<sup>b</sup>M. P. Council of Science and Technology, Vigyan Bhawan, Bhopal (M.P.)

**KEY WORDS:** Spatial Variability, Soil Properties, semi -variogram, Sanwer, Kriging, GPS

### ABSTRACT:

A study was conducted to explore the spatial variability of major soil nutrients in a soybean grown region of *Malwa* plateau. From the study area, one hundred sixty two surface soil samples were collected by a random sampling strategy using GPS. Then soil physico-chemical properties i.e., pH, EC, organic carbon, soil available nutrients (N, P, K, S and Zn) were measured in laboratory. After data normalization, classical and geo-statistical analyses were used to describe soil properties and spatial correlation of soil characteristics. Spatial variability of soil physico-chemical properties was quantified through semi-variogram analysis and the respective surface maps were prepared through ordinary Kriging. Exponential model fits well with experimental semi-variogram of pH, EC, OC, available N, P, K, S and Zn. pH, EC, OC, N, P, and K has displayed moderate spatial dependence whereas S and Zn showed weak spatial dependence. Cross validation of kriged map shows that spatial prediction of soil nutrients using semi-variogram parameters is better than assuming mean of observed value for any un-sampled location. Therefore it is a suitable alternative method for accurate estimation of chemical properties of soil in un-sampled positions as compared to direct measurement which has time and costs concerned.

### 1. INTRODUCTION

Soil macronutrients are essential to plants growth; maintain ecosystems and high crop yields. However, imbalance fertilization, deteriorate the precious soil environment particularly N and P can be potentially hazardous to water resources when their available components in soils are excessive, because available macronutrients can be transported off site in runoff due to rain or irrigation (Smith *et al.*, 1998, Phupaibul *et al.*, 2004 and Ju *et al.*, 2007) and subsequently degrades the fertility of soil and reduced the productivity.

Several studies have documented that soil properties vary across agricultural fields, causing spatial variability in crop yields. Therefore, their proper management is necessary to avoid deteriorating the environment while meeting the requirement of high crop productivity and farmer must be advised to use balanced of fertilizers/manures, special soil

amendment (if any) and accordingly adopt suitable cropping pattern. Hence it is necessary to evaluate the fertility status of the soil and promote the recommendations of soil test for balanced nutrition to maintain soil health. The application of parametric statistics is inadequate for analysis of spatially dependant variables because, they assume that measured observations are independent in spite of their distribution in space. Now days, Geographic information systems (GIS), as new technology, and Geo-statistics provide a tool for improving sampling design by utilizing the spatial dependence of soil properties within a sampling region and useful to illustrate the spatial interrelationship of soil data which reduces error, biasness and increase the accuracy of data for interpolation. Therefore, the present study has been planned to quantify the spatial variability of soils in Sanwer of Indore district of M.P.

### 2. MATERIALS AND METHODS

#### 2.1 Study Area

Geographically, the Sanwer is located between 22°55'48.90" N latitudes and 75°51'10.80"E longitudes in Indore district, M.P having temperature range of 25°C to 44°C in summer and 10°C to 31°C in winter. The total area of the block is 694 sq km. Based on soil taxonomy (USDA, 2010), this region has Vertisols and associated soil orders. These soils are montmorillonitic, neutral to slight alkaline and having high swell shrink potential.

Soybean (*Glycine max*) is an important oilseed grown as rain fed crop (80% area) during *kharif* season in the semi-arid *Malwa* region of Madhya Pradesh. The farmers usually apply a lower dose of fertilizer for soybean in study area impose the assessment and delineation of the extent of nutrients status and their variability in soils of Sanwer tehsil, so that a picture regarding depletion and build up may be obtained.

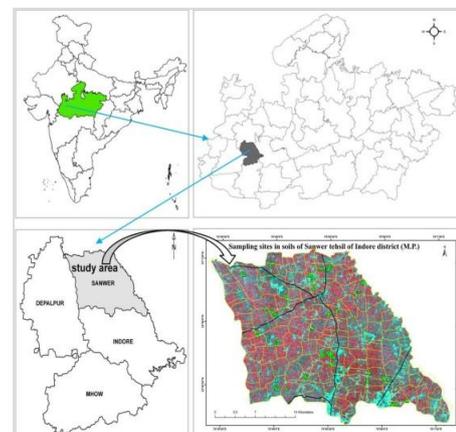


Figure-1 Location map of study area

\*Corresponding author: bairagigd@gmail.com

## 2.2 Soil sampling and their chemical analysis

In the present study, one hundred sixty two topsoil samples were collected from soybean-wheat cropping system in 2011-12. The sampling scheme considered the distributional uniformity of samples and also ensured that all samples were located in soybean-wheat fields (Figure-1). All samples were taken in fall after harvest and before next cropping season to avoid the effect of fertilization during crop cultivation. For each soil sample, soils at 6– 8 points within a small area of approximately 0.01 ha were collected from surface layers (0–15 cm) and then mixed and delivered to a laboratory for analysis. Exact sample locations were recorded using a hand-held global position system. All samples were air-dried at room temperature (20–22°C), ground using wooden mortar and pestle, and sieved into soil particles less than 2 mm.

Soil pH was measured using a pH meter (Systronic pH meter) with a soil/water ratio of 1: 2.5 (Jackson, 1973). Electrical conductivity was determined in 1:2.5 soil: water supernatant solution with the help of conductivity -bridge (Jackson, 1973). Soil organic carbon was estimated using the Walkley and Black wet oxidation method (Jackson, 1973). Available nitrogen was estimated by alkaline permanganate

## 2.3 Geo-statistical analysis

The presence of a spatial structure where observations close to each other are more alike than those that are far apart (spatial autocorrelation) is a prerequisite to the application of geo-statistics. The experimental variogram measures the average degree of dissimilarity between unsampled values and a nearby data value (Deutsch and Journel, 1998) and thus can depict autocorrelation at various distances. The value of the experimental variogram for a separation distance of  $h$  (referred to as the lag) is half the average squared difference between the value at  $z(x_i)$  and the value at  $z(x_i+h)$  (Lark, 2000; Robinson and Metternicht, 2006):

method (Subbiah and Asija, 1956). The available Phosphorus was estimated by 0.5 M sodium bicarbonate method as described by (Olsen *et al.*, 1954) and the P concentration was quantified using spectrophotometer. Available Potassium was extracted with Neutral 1 N ammonium acetate as described by Jackson (1973) and determined by flame photometer. Available S (AS) was extracted using 0.15%  $\text{CaCl}_2$  and then determined by the turbidimetric method Chesnin and Yien (1951). Available Fe, Available Mn, Available Cu, and Available Zn were extracted with diethylenetriaminepentaacetic acid (DTPA), and the extracted Fe, Mn, Cu, and Zn were determined with flame atomic absorption spectrometry by Lindsay and Norvell. (1978).

The nutrient index (NI) values for available nutrients present in the soils were calculated utilizing the formula suggested by Parker *et al.* (1951) and classified this index as low (<1.67), medium (1.67 to 2.33) and high (>2.33).

$$\text{NI} = \frac{(\text{Nl} \times 1) + (\text{Nm} \times 2) + (\text{Nh} \times 3)}{\text{Nt}}$$

Where: N1, Nm and Nh are the number of soil samples falling in low, medium and high categories for nutrient status and are given weightage of 1, 2 and 3, respectively. Nt is the total number of samples.

$$\gamma(h) = \frac{1}{2N(h)} \sum_{i=1}^{N(h)} [z(x_i + h) - z(x_i)]^2$$

Where:  $N(h)$  is the number of data pairs within a given class of distance and direction. If the values at  $z(x_i)$  and  $z(x_i+h)$  are auto correlated the result of Eq.(1) will be small, relative to an uncorrelated pair of points. From analysis of the experimental variogram, a suitable model is then fitted, usually by weighted least squares, and the parameters (e.g. range, nugget and sill) are then used in the Kriging procedure.

## 3. RESULTS AND DISCUSSIONS

### 3.1 Descriptive Statistics of Soil Properties

The mean values for pH, EC, OC content were 7.57, 0.45  $\text{dSm}^{-1}$  and 6.99  $\text{g kg}^{-1}$  with a range of 6.53-8.52, 0.08 to 1.24  $\text{dS m}^{-1}$  and 2.6 to 9.95  $\text{g kg}^{-1}$ , respectively. The available N, P, and K varied from 167.98 to 393.66  $\text{kg ha}^{-1}$ , 7.77 to 39.32  $\text{kg ha}^{-1}$  and 104.56 to 851.6  $\text{kg ha}^{-1}$  with mean value of 273.78  $\text{kg ha}^{-1}$ , 19.38  $\text{kg ha}^{-1}$  and 461.35  $\text{kg ha}^{-1}$ , respectively. The S varied from 6 to 20  $\text{mg kg}^{-1}$  with a mean value of 10.8  $\text{mg kg}^{-1}$ . The available micronutrients Zn, Cu, Fe and Mn varied from 0.17 to 1.58, 0.62 to 7.8, 1.96 to 14.76 and 1.06 to 20.92  $\text{mg kg}^{-1}$  with mean values of 0.68, 3.01, 6.84 and 10.14  $\text{mg kg}^{-1}$ .

The coefficient of variation, which is the ratio of the standard deviation to mean expressed as a percentage is a useful measure of overall variability. The EC had the largest variation (CV = 68.89 per cent) followed by OC (CV = 23.58 per cent) while pH was found least variable (CV = 6.37 per cent). The macronutrients are i.e., NPK, the available P had the highest variability (CV=44.43 per cent) and followed by avail. K (CV=42.73 per cent) whereas N had the lowest variability (CV= 20.16). The S was found to be moderately variable (CV = 31.44 per cent). Among the micronutrients, the Cu was found to be highly variable (CV = 61.81 per cent), followed by Zn (CV=49.59 per cent) and Mn (CV= 45.49 percent) while Fe

was found least variable (CV= 38.13 percent). According to the classification of Hillel (1980), pH showed low variability (<10%) and other variables except Zn showed medium variability (10% - 100%) whereas available Zn had a high variability (>100%). The range of CV for the area suggested different degrees of heterogeneity among the properties studied. All the macronutrients were found to be moderately variable with ranging from 20.26-44.43 per cent. All the micronutrients were highly variable with CV ranging from 38.13–61.81 per cent.

The normality of data was tested by Kolmogorov-Smirnov (K-S) method (P-value > 0.05) and a result of soil parameters is presented in Table 1. With due attention to the levels of skewness for these parameters were normal. The skewness and kurtosis coefficients are zero for a normally distributed random variable. If the data distributions are largely deviated from a normal distribution, data transformations are often performed in order to reduce the influence of extreme values on spatial analysis (Webster and Oliver, 2001). However, considering the observed skewness coefficient values data were not transformed. Among the soil fertility parameters, EC, available P, available S, Zn, Cu and Fe were found not

normally distributed due to higher value of skewness and kurtosis.

Table-1 Statistical overview for physico-chemical properties of soils of study area (N=162)

Parameters	Range	Minimum	Maximum	Mean	S.D.	Skewness	Kurtosis	CV%
pH	1.99	6.53	8.52	7.57	0.48	-0.57	-0.38	6.37
EC dSm <sup>-1</sup>	1.16	0.08	1.24	0.45	0.31	<b>1.01</b>	-0.14	68.89
lnEC dSm <sup>-1</sup>	2.735	-2.52	0.215	-1.031	0.689	0.03	2.15	-66.83
OC gkg <sup>-1</sup>	7.35	2.6	9.95	6.99	1.65	-0.44	0.06	23.58
Avail. N kg ha <sup>-1</sup>	225.68	167.98	393.66	273.78	55.46	0.07	-0.68	20.26
Avail. P kg ha <sup>-1</sup>	31.55	7.77	39.32	19.38	8.61	<b>0.85</b>	-0.21	44.43
ln Avail. P kg ha <sup>-1</sup>	1.62	2.05	3.67	2.87	0.43	0.171	2.13	14.98
Avail. K kg ha <sup>-1</sup>	747.04	104.56	851.6	461.35	197.12	0.03	-0.98	42.73
S mgkg <sup>-1</sup>	14	6	20	10.8	3.4	<b>0.55</b>	-0.51	31.44
ln S mgkg <sup>-1</sup>	1.20	1.79	2.99	2.33	0.32	0.05	2.05	13.65
Zn mgkg <sup>-1</sup>	1.41	0.17	1.58	0.68	0.34	<b>0.71</b>	-0.39	49.59
ln Zn mgkg <sup>-1</sup>	2.227	-1.77	0.457	-0.51	0.515	-0.197	2.4	-100.98
Cu mgkg <sup>-1</sup>	7.18	0.62	7.8	3.01	1.86	<b>0.87</b>	-0.17	61.81
ln Cu mgkg <sup>-1</sup>	2.52	-0.47	2.05	0.907	0.643	-0.103	2.12	70.89
Fe mgkg <sup>-1</sup>	12.8	1.96	14.76	6.84	2.61	<b>0.6</b>	0.16	38.13
ln Fe mgkg <sup>-1</sup>	2.016	0.674	2.69	1.84	0.398	-0.339	2.77	21.63
Mn mgkg <sup>-1</sup>	19.85	1.06	20.92	10.14	4.61	0.17	-0.46	45.49

### 3.2 Soil fertility status using classical statistics

On the basis of the ratings suggested by Subbiah and Asija (1956), 37.65 per cent samples were found to be low while 62.3 were medium and none of samples were high. On the basis of the limits suggested by Muhr *et al.* (1963), only 9.26 per cent soil samples were found low, 56.2 were found in medium and rest of high (<20 P<sub>2</sub> O<sub>5</sub> kg ha<sup>-1</sup>). This might be due to the presence of phosphorus in organic forms and after

decomposition of organic matter as humus is formed which forms complex with Al and Fe and that is a protective cover for P fixation with Al and Fe thus reduce phosphorus adsorption/ Phosphate fixation (Tisdale *et al.* 1997). According to Muhr *et al.* (1963) 16.67 samples were low, 25.3 percent samples were found in medium and remaining (58.0) were high in potassium content.

The limits of micronutrients (mg kg<sup>-1</sup>) were used for various categories (low, medium and high) as suggested by (Singh *et al.*, 2007) are given as under:

Micronutrients	Low	Medium	High
Zn	<0.6	0.6-1.2	>1.2
Cu	<0.2	0.2-0.4	>0.4
Fe	<4.5	4.5-9.0	>9.0
Mn	<2.0	2.0-4.0	>4.0
B	<0.1	0.1-0.60	>0.60
N.I.	<1.67	1.67-2.33	>2.33

Result revealed the soils were low in organic carbon content and for these soils available micronutrients status using 0.6, 0.2 and 4.5 mg kg<sup>-1</sup> soil as the critical limit for DTPA-extractable for Zn, Cu, and Fe (Follet and Lindsay 1970) 51.85 percent soil samples were found to be deficient in available Zn. The data from the table showed the 1.0 mg kg<sup>-1</sup> and 0.2 mg kg<sup>-1</sup> soil as the critical limit for DTPA-extractable Mn and Cu respectively (Follet and Lindsay 1970), none of the soils samples were tested low in Mn and Cu content respectively. All soil samples found in sufficient quantity of available Cu content. Similar reported by Mehra *et al.* (2005). Considering 10 mg kg<sup>-1</sup> as the threshold value (Balanagoudar &

Satyanarayan, 1990), 45.68 and 54.3 percent samples were found to be deficient and sufficient respectively in Sanwer tehsil. Results are supported by Rathore *et al.* (1995).

The nutrient index (NI) values for available nutrients present in the soils were calculated utilizing the formula suggested by Parker *et al.* (1951) and classified this index as low (<1.67), medium (1.67 to 2.33) and high (>2.33). Considering soil nutrient index (Table 2) soils of Sanwer were found of high status in respect of K, Mn and Cu, medium status in respect of OC, P and Fe S, while low fertility status in case of N,S and Zn.

Table 2 Percent sample deficiency and NI generated using classical statistics of (N=162)

Nutrients parameters		Percent samples under different categories			NI	NI class
		LOW	MEDIUM	HIGH		
OC		11.73	35.2	53.1	2.23	Medium
Available macronutrients	N	37.65	62.3	0.0	1.62	Low
	P	9.26	56.2	34.6	2.25	Medium
	K	16.67	25.3	58.0	2.41	High
Secondary nutrient (S)		45.68	0.0	54.3	1.54	Low
Micronutrients	Zn	51.85	38.3	9.9	1.58	Low
	Cu	0	0.0	100.0	3.00	High
	Fe	22.22	57.4	20.4	1.98	Medium
	Mn	0	10.5	89.5	2.90	High

### 3.3 Relationship with micronutrients

Correlation studies showed the significant positive correlation between organic carbon and available N, K were found with having values of  $r=0.91$  and  $r= 0.29$  respectively. The

relationship between availability of N, Fe with available K showing value  $r=.199^*$   $r =0.188^*$  also exist.

Table 3 Pearson's correlation coefficients between major, S, micronutrients and other soil properties

Parameters	pH	EC	OC	Ava N	Ava P	Ava K	Zn	Cu	Fe	Mn
EC	-.205**	1								
OC	0.052	-0.031	1							
Ava N	0.04	0.007	<b>.905**</b>	1						
Ava P	-0.007	<b>-.271**</b>	0.144	0.149	1					
Ava K	.313**	0.115	<b>.291**</b>	<b>.199*</b>	-0.069	1				
Zn	-0.122	0.146	0.034	0.025	-0.152	0.001	1			
Cu	0.082	-0.063	-0.018	-0.022	-0.042	0.071	0.087	1		
Fe	0.065	0.054	0.0	-0.045	-0.052	<b>.188*</b>	0.028	-0.005	1	
Mn	-0.006	0.007	0.025	-0.018	-0.074	-0.039	0.011	-0.041	<b>-.371**</b>	1
S	-0.112	0.06	-0.037	-0.008	-0.04	0.021	0.127	-0.012	0.073	0.007

\*\* . Correlation is significant at the 0.01 level (2-tailed).

\* . Correlation is significant at the 0.05 level (2-tailed).

The relationship between available Mn and Fe were found to be negative and statistically significant, ( $r=-0.371^{**}$ ), Correlation between EC and pH, available P were found

negative significant ( $r=-0.205^{**}$ ) and ( $r=-0.271^{**}$ ) whereas available K and pH were found positive and significant, ( $r=0.313^{**}$ ).

### 3.4 Spatial variability

Geo-statistical methods were used to analyze the spatial correlation structures of the available contents of N, P, K and S and micronutrients in soil and spatially estimate their values at unsampled locations. Because Kriging assumes the normal distribution for each estimated variable, it is necessary to check whether the available contents of N, P, K and S and micronutrients in soil samples are approximately normally distributed or not. The first step in using of Kriging methods is to check the presence of spatial structure among data by variogram analysis. The distribution of data should be normal for the parameter estimation, and the K-S test was used to examine the distribution of the data. From Table 1, one can see that the skewness and kurtosis indices of all the soil properties are close to the standard value of 0, but those of S, EC, and Zn are not. In this study, natural logarithmic transformation was used to reduce the skewness of the data distributions of P and K. The histograms of N, P, and K

data used for geo-statistical interpolation are shown in Figure 2. One can see that they tend to be normally distributed. The information generated with variogram was used to calculate sample weighing factors for spatial interpolation by ordinary Kriging procedures. Ordinary Kriging was chosen to create the spatial distribution maps of soil N, P, and K contents, with the maximum search radius being set to the autocorrelation range of the corresponding variable. The best model for fitting on experimental variogram was selected based on less RSS value. The exponential model was selected from standard models that are available to fit experimental semi-variograms based on more favorable weighted residual mean squares, and visual fit to the data at short lags. Therefore, we recognized the exponential and spherical model to be suitable for estimation of soil properties. The variograms of studied soil parameters are shown in Table 3, Figures 2 and 3.

In this study, no apparent anisotropy was found for any studied variable through experimental variograms. So, all experimental variograms were in isotropic form and were fitted using basic math models, such as Spherical, exponential, Gaussian and linear based on the values of weighted residual sums of squares, regression coefficient ( $r^2$ ) and relative spatial structure indicator (Nugget/Sill) that indicated spatial dependency for Kriging interpolation. The ratio of nugget variance to sill expressed in percentages ( $C_0/C+C_0$ ) can be regarded as a criterion for classifying the spatial dependence of the soil parameters. If this ratio is less than 25%, then the variable has strong spatial dependence (Shi *et al.*, 2005). The pH, EC, OC, N, P, K, Zn and Cu showed the values between 25% and 75% considered as moderate spatial dependence, and the values greater than 75% have weak spatial dependence as shown in table, S Fe and Mn have moderate spatial structure. A filled contour map (prediction map) and a relevant prediction standard error map were created for soil properties using the ArcGIS Geo-statistics tool.

ln (AP), K, (Zn), ln(Cu), and ln(Fe) and spherical models for Mn. The nugget/sill ratios of variogram models for pH, EC, OC, N, P, K, Zn and Cu all fall between 25% and 75%, which exhibit moderate spatial dependency. Usually, a strong spatial dependence of soil properties can be attributed to intrinsic factors, and a weak spatial dependence can be attributed to extrinsic factors (Cambardella *et al.*, 1994). The correlation range measures the spatial separate distances within which data are auto correlated (Cahn *et al.*, 1994). The approximate correlation ranges for pH, EC, OC, AN, ln(AP), K,S, (Zn), ln(Cu), ln (Fe) and Mn are 3.112 km, 0.783 km, 2.14 km, 0.706 km, 3.13 km, 1.29 km, 16.57 km, 4.74 km, 1.53 km, 14.64 km, and 0.602 km, respectively. Apparently, pH, OC, AK, S, (Zn), ln(Cu) and ln(Fe) are auto correlated in longer ranges than AN, EC, and Mn are. This result is consistent with their CV values. This may imply that AP, and AK contents are more sensitive to extrinsic factors such as fertilization. To map the spatial distributions of pH, EC, OC, AN, ln (AP), K,S, (Zn), ln(Cu), ln(Fe) and Mn, we used ordinary Kriging to interpolate their respective sample data (Figure 4).

Experimental variograms and fitted models for are presented in **Figure 3**. Exponential models are chosen for pH, EC, OC, AN,

Table 4 Characteristic of calculated semivariogram of spatial soil fertility using ordinary Kriging

Soil property	Semi-varigarm characteristic					
	Range (m)	Nugget (C0)	Partial Sill (C)	Sill (C <sub>0</sub> +C)	NS ratio	Spatial dependence
pH	3112.0	0.127	0.120	0.247	0.51	Moderate
ln EC	783.5	0.248	0.159	0.407	0.61	Moderate
OC	2143.9	1.520	1.377	2.896	0.52	Moderate
N	706.7	1718.10	581.46	2299.6	0.75	Moderate
lnP	3130.5	0.132	0.059	0.192	0.69	Moderate
K	1295.1	15296.0	23788.0	39084.0	0.39	Moderate
lnS	16572.1	0.092	0.006	0.098	0.94	Weak
lnZn	4743.0	0.208	0.037	0.245	0.85	Weak
lnCu	1531.1	0.267	0.190	0.458	0.58	Moderate
lnFe	14640.0	0.157	0.003	0.160	0.98	weak
Mn	602.0	22.089	0.118	22.207	0.99	weak

The Mean Absolute Error (MAE) was found to be zero for pH, EC, OC, P, S, Zn, Cu, Fe and Mn. Highest Mean Squared Error (MSE) was observed for K followed by N, P, Mn, S, Fe, Cu, OC, pH, Zn and EC. The goodness of fit (G) values was positive and highest G value was observed for K followed

by pH. Lower MSE value was observed for organic carbon, EC and pH and Zn. Higher G value was observed for pH compared to EC N, P organic carbon and Cu. The G values for S, Zn, Fe and Mn were negative. Higher G values were observed for available N, P compared to EC, OC and available Cu.

Table-5 Evaluation performance of Krigged map of soil properties through cross-validation

parameters	MAE	MSE	G
pH	0.0	0.17	24.47
lnEC	0.0	0.09	7.78
OC	0.0	2.52	6.83
N	-0.6	2585.58	15.40
lnP	0.0	62.13	15.67
K	-2.9	27338.03	29.20
lnS	0.0	11.59	-1.092
lnZn	0.0	0.11	-0.063
lnCu	0.0	3.24	6.12
lnFe	0.0	7.23	-6.99
Mn	0.0	24.98	-18.15

Among the two different theoretical models tested, the exponential model was found best fit for soil properties including pH, EC, OC, N, P, K, S, Zn, Cu and Fe whereas, spherical model was found as the best fit for Mn. In order to identify the spatial distribution patterns of soil properties in the surveyed area, it is necessary to present soil properties data in

the form of a map. For this purpose, spatial distribution maps of all fertility parameters in soil obtained by ordinary Kriging based on exponential model for surface (0 – 15cm) soil in Sanwer tehsil are presented in figure 3. The results also supported by Reza *et al.*, 2010.

#### 4.0 CONCLUSIONS

Spatial variability of soil fertility was quantified through semivariogram analysis and interpolated through ordinary Kriging using best fit exponential model. Based on prediction and error maps, it was realized that all soil nutrients has no toxic status and that the application of fertilizer will

improve crop yields in the study area. Cross validation of kriged maps shows spatial prediction of soil properties with reasonable accuracy. It was concluded that geo-statistical techniques are applicable to investigate the spatial variability of soil properties in the study area.

#### REFERENCE

- Cahn M. D., J. W. Hummel, and Brouer, B. H. 1994. Spatial analysis of soil fertility for site-specific crop management. *Soil Science Society of America Journal*, 58, (4), pp. 1240–1248.
- Cambardella, C.A., Moorman, T.B., Novak, J.M., Parkin, T.B., Karlen, D.L., Turco, R.F. and Konopka, A.E., 1994. Field scale variability of soil properties in Central Iowa soils. *Soil Science Society of America Journal*, 58, 1501–1511.
- Follet R. H. and Lindsay, W.L. 1970. Profile distribution of Zn, Cu, Fe and Mn in Colorado State University Experimental Station. Fort Collin Co. Technical Bulletin 110.
- Goovaerts, P. 1997. *Geo-statistics for Natural Resources Evaluation*, Oxford University Press, New York, NY, USA.
- Hillel, D. (1980). *Fundamentals of Soil Physics*. Academic Press, Inc., New York
- Ju X. T., C. L. Kou, P. Christie, Z. X. Dou, and F. S. Zhang, 2007. Changes in the soil environment from excessive application of fertilizers and manures to two contrasting intensive cropping systems on the North China Plain. *Environmental Pollution*, vol. 145 (2), pp. 497–506.
- Lindsay, W.L. and W.A. Norvell. 1978. Development of a DTPA soil test for Zn, Fe, Mn and Cu. *Soil Sci. Soc. Am. J.* 42: 421-428.
- Muhr G.R., Datta N.P., Shankara Subraney N., Dever F., Lecy V.K. and Donahue R.R., 1963. Soil Testing in India. USAID Mission to India.
- Parker, F.W., Nelson, W.L., Winter Eric and Miller, L.E. 1951. The broad interpretation of soil test informations. *Agronomy Journal* 43, 105-112.
- Phupaibul, P., C. Chitbuntanorm, N. Chinoim, P. Kangyawongha, and T. Matoh, 2004. Phosphorus accumulation in soils and nitrate contamination in underground water under export-oriented asparagus farming in Nong Ngu Lauem Village, Nakhon Pathom Province, Thailand. *Soil Science and Plant Nutrition*, 50(3), pp. 385–393.
- Reza, S.K., Sarkar, D., Baruah, U. and Das, T.H., 2010. Evaluation and comparison of ordinary Kriging and inverse distance weighing methods for prediction of spatial variability of some chemical parameters of Dhalai district, Tripura. *Agropedology*, 20, 38–48.
- Shi Z, Li Y, Makeschine F, Wang RC 2005. Assessment of temporal and spatial variability of soil salinity in a coastal saline field. *Environ Geo.* 48:171-178.
- Singh, D., Chhonkar, P. K. and Dwivedi, B. S. 2007. *Manual on Soil, Plant and Water Analysis*. Westville Publishing House, New Delhi.
- Smith V. H., G. D. Tilman, and J. C. Nekola 1998. Eutrophication: impacts of excess nutrient inputs on freshwater, marine, and terrestrial ecosystems. *Environmental Pollution*, 100 (1– 3), pp. 179–196.
- Subbiah, B. V. and Asija G. L. 1956. A rapid procedure for the determination of available nitrogen in soils. *Curr. Sci.*, 25: 259-260.
- Tisdale, I. S., I. W. Nelson, D. J. Beaton and Havlin I. J. 1997. *Soil Fertility and Fertilizers*. 5th ed. Prentice Hall of India.
- Webster, R. and Oliver, M.A. 2001. *Statistical Methods in Soil Science and Resource Survey*. Oxford University Press, New York, NY.

Figure 2: Histograms of pH (a), EC (b), OC (c), AN (d), AP (e) and AK (f),

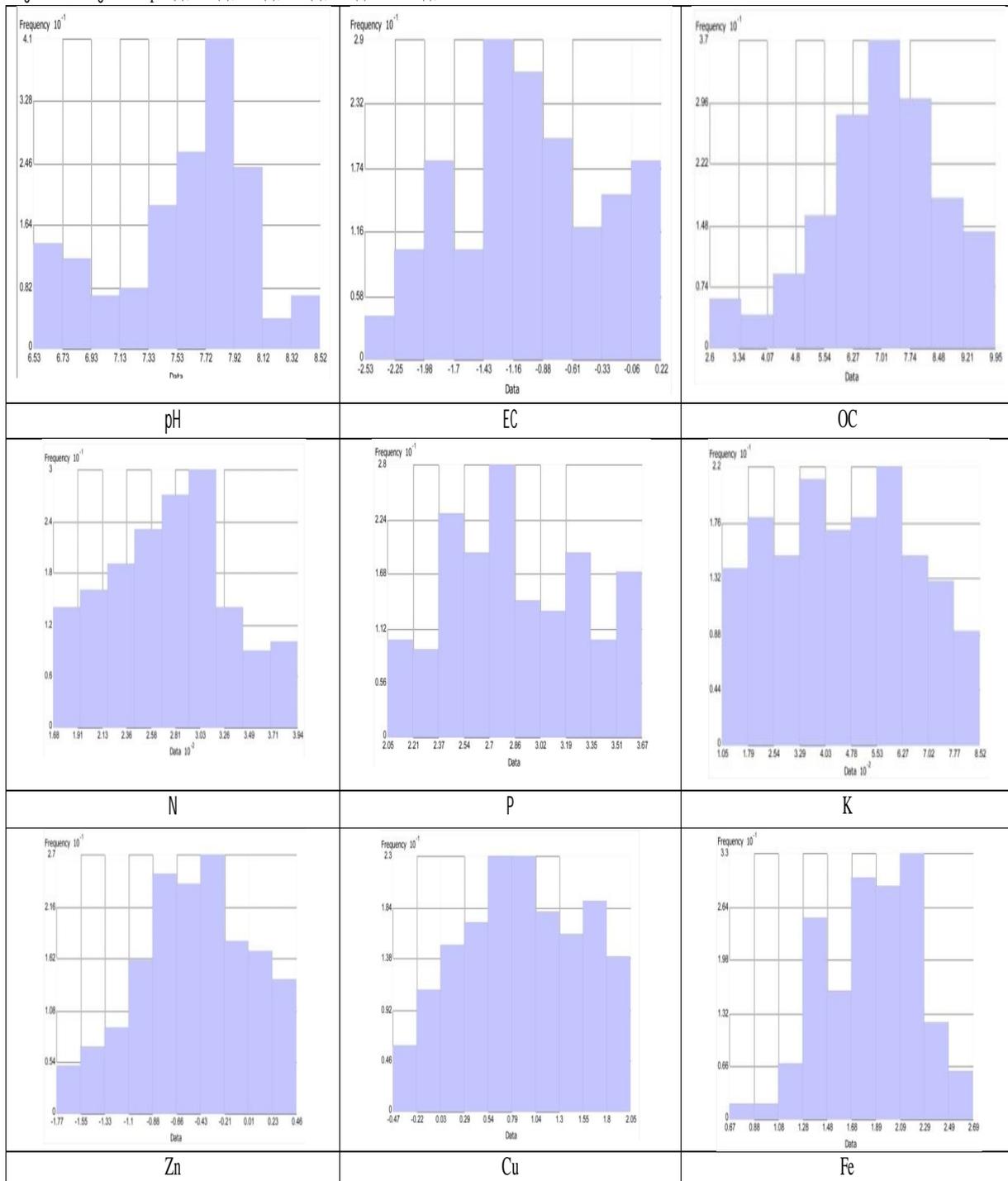


Figure 3: Variograms of pH (a), EC (b), OC (c), AN (d), AP (e) and AK (f)

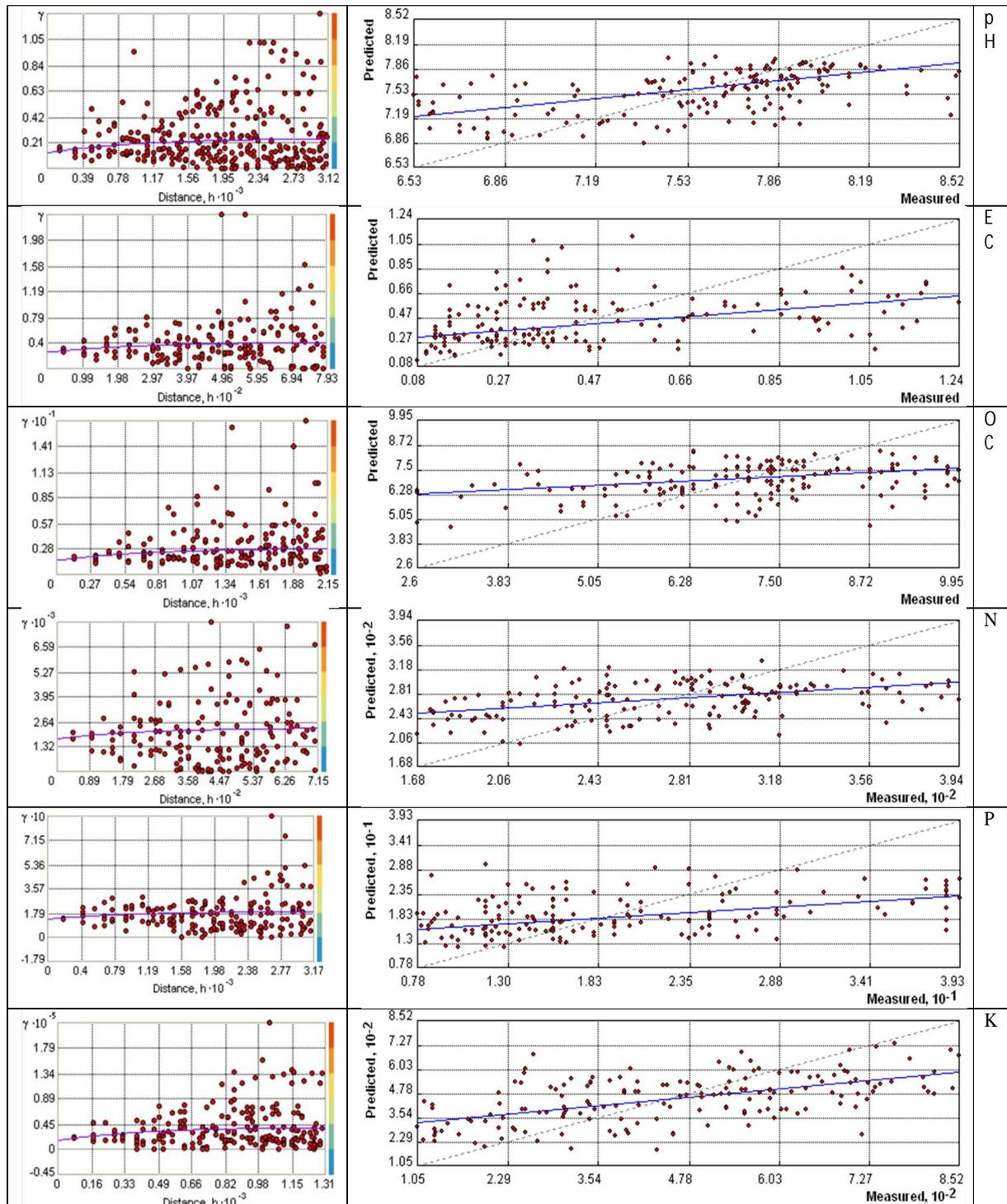


Figure 4: Spatial distribution maps of pH (a), EC (b), OC (c), AN (d), AP (e), AK (f), AS (g), Zn(h),Cu(i), Fe(j), Mn (k) and Multi macronutrients map (l), interpolated by ordinary Kriging.

